

Evidence Synthesis / Meta-Analysis

Session 2, Lecture 4: Meta-Analysis with Binary Outcome

Guido Knapp¹, Gerta Rücker², Guido Schwarzer²

¹ Department of Statistics, TU Dortmund University

² Center for Medical Biometry and Medical Informatics, University of Freiburg

sc@imbi.uni-freiburg.de

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Example: Aggressive Non-Hodgkin Lymphoma

Greb et al. (2008), Cochrane Database Syst Rev 1, CD004024:

- ▶ Cochrane Review including 15 randomised controlled trials (RCTs)
- ▶ Adult patients with aggressive non-Hodgkin lymphoma
- ▶ First line treatment with high-dose chemotherapy (HDCT) versus conventional chemotherapy
- ▶ Primary outcome:
Overall survival (14 RCTs, 2444 patients)
- ▶ **Secondary outcome:**
Complete response (14 RCTs, 2126 patients)

Overview Lecture 4

- ▶ Standard methods of meta-analysis with binary outcome
 - ▶ Fixed effect methods (Inverse variance, Mantel-Haenszel, Peto)
 - ▶ Random effects method (Inverse variance)
- ▶ Peculiarities of sparse binary data
- ▶ Generalised linear mixed model
 - ▶ Conditional model, exact likelihood (Hypergeometric-Normal model)
 - ▶ Conditional model, approximate likelihood (Binomial-Normal model)

Aggressive Non-Hodgkin Lymphoma – Complete Response

Study	HDCT		Control	
	Events	Total	Events	Total
De Souza	14	28	10	26
Gianni	46	48	35	50
Gisselbrecht	119	189	116	181
Intragumtornchai	10	23	9	25
Kaiser	110	158	97	154
Kluin-Nelemans	67	98	56	96
Martelli 1996	3	22	4	27
Martelli 2003	57	75	51	75
Milpied	74	98	56	99
Rodriguez 2003	39	55	30	53
Santini 1998	46	63	34	61
Santini-2	80	117	71	106
Verdonck	25	38	26	35
Vitolo	35	60	46	66

	CR		no CR			
HDCT	74	(a)	24	(b)	98	(a + b = n _T)
Control	56	(c)	43	(d)	99	(c + d = n _C)
	130	(a + c)	67	(b + d)	197	(n)

```
mil <- metabin(crHDCT, nHDCT, crControl, nControl,
  data = cr, subset = study == "Milpied",
  sm = "OR")
```

```
# Print odds ratio for Milpied study
round(exp(mil$TE), 2)

## [1] 2.37

# Print risk ratio
round(exp(update(mil, sm = "RR")$TE), 2)

## [1] 1.33

# Print risk difference
round(update(mil, sm = "RD")$TE, 2)

## [1] 0.19
```

Let

- ▶ p_T : Experimental event probability $\hat{p}_T = a/(a + b)$
- ▶ p_C : Control event probability $\hat{p}_C = c/(c + d)$

Risk Ratio ϕ :

$$\phi = \frac{p_T}{p_C} \quad \hat{\phi} = \frac{\hat{p}_T}{\hat{p}_C}$$

Odds Ratio ψ :

$$\psi = \frac{\left(\frac{p_T}{1 - p_T}\right)}{\left(\frac{p_C}{1 - p_C}\right)} = \phi \times \frac{1 - p_C}{1 - p_T} \quad \hat{\psi} = \frac{a d}{b c} \quad (1)$$

Risk Difference η :

$$\eta = p_T - p_C \quad \hat{\eta} = \hat{p}_T - \hat{p}_C$$

```
# Calls R function rma.uni (Random effects Meta-Analysis - UNivariate)
mil4 <- rma(ai = crHDCT, n1i = nHDCT, ci = crControl, n2i = nControl,
  data = cr, subset = study == "Milpied",
  measure = "OR")
```

```
round(exp(mil4$b), 2)

##          [,1]
## intrcpt 2.37

round(exp(update(mil4, measure = "RR")$b), 2)

##          [,1]
## intrcpt 1.33

round(update(mil4, measure = "RD")$b, 2)

##          [,1]
## intrcpt 0.19
```

Binary Effect Measures – Confidence Interval

Large sample variance estimates (Fleiss, 1993):

$$\begin{aligned}\widehat{\text{Var}}(\log \hat{\phi}) &= \frac{1}{a} + \frac{1}{c} - \frac{1}{a+b} - \frac{1}{c+d} \\ \widehat{\text{Var}}(\log \hat{\psi}) &= \frac{1}{a} + \frac{1}{b} + \frac{1}{c} + \frac{1}{d} \\ \widehat{\text{Var}}(\hat{\eta}) &= \frac{ab}{(a+b)^3} + \frac{cd}{(c+d)^3}\end{aligned}\quad (2)$$

$(1 - \alpha)$ -confidence interval (on log scale for risk ratio and odds ratio):

$$\hat{\theta} \pm z_{1-\frac{\alpha}{2}} \text{S.E.}(\hat{\theta})$$

with standard error $\text{S.E.}(\hat{\theta}) = \sqrt{\widehat{\text{Var}}(\hat{\theta})}$.

Binary Effect Measures – Confidence Interval

```
# Print confidence interval for odds ratio (R package meta)
print(mil, digits = 2)
```

```
##      OR      95%-CI      z  p-value
## 2.37 [1.29; 4.35] 2.78  0.0055
##
## Details:
## - Inverse variance method
```

```
# Print confidence interval for log odds ratio (R package meta)
print(mil, digits = 2, backtransf = FALSE)
```

```
## logOR      95%-CI      z  p-value
## 0.86 [0.25; 1.47] 2.78  0.0055
##
## Details:
## - Inverse variance method
```

Binary Effect Measures – Confidence Interval

Large sample variance estimates (Fleiss, 1993):

$$\begin{aligned}\widehat{\text{Var}}(\log \hat{\phi}) &= \frac{1}{a+0.5} + \frac{1}{c+0.5} - \frac{1}{a+b+0.5} - \frac{1}{c+d+0.5} \\ \widehat{\text{Var}}(\log \hat{\psi}) &= \frac{1}{a+0.5} + \frac{1}{b+0.5} + \frac{1}{c+0.5} + \frac{1}{d+0.5} \\ \widehat{\text{Var}}(\hat{\eta}) &= \frac{(a+0.5)(b+0.5)}{(a+b+1)^3} + \frac{(c+0.5)(d+0.5)}{(c+d+1)^3}\end{aligned}$$

Add 0.5 if any cell counts are zero (Gart and Zweifel, 1967; Pettigrew et al., 1986)

Default in **metabin** (argument **incr**) and **rma** (argument **add**)

Binary Effect Measures – Confidence Interval

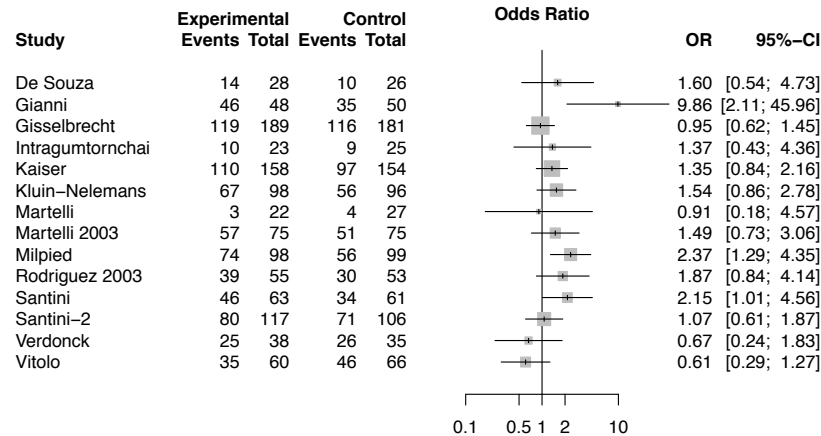
```
print(mil4, digits = 2) # log odds ratio (R package metafor)

##
## Fixed-Effects Model (k = 1)
##
## Test for Heterogeneity:
## Q(df = 0) = 0.00, p-val = 1.00
##
## Model Results:
##
## estimate      se      zval      pval      ci.lb      ci.ub      **
##      0.86      0.31      2.78      <.01      0.25      1.47
##
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

print(predict(mil4, transf = exp), digits = 2) # odds ratio

## pred ci.lb ci.ub
## 2.37 1.29 4.35
```

Forest Plot – CR



Naive Pooling – Fictitious Example

		CR	no CR	\hat{p}_T	\hat{p}_C	\widehat{RR} [95%-CI]
Study 1	HDCT	4	56	6.7%	7.3%	0.91 [0.30; 2.74]
	Control	11	139			
Study 2	HDCT	40	140	22.2%	24.0%	0.93 [0.53; 1.63]
	Control	12	38			
Study 1&2	HDCT	44	196	18.3%	11.5%	1.59 [1.00; 2.55]
	Control	23	177			
Appropriate meta-analysis						0.92 [0.56; 1.52]

Inverse Variance Method – Odds ratio – Definition

Overall odds ratio $\hat{\psi}_{IV}$ (Fleiss, 1993):

$$\hat{\psi}_{IV} = \exp \left(\frac{\sum_{k=1}^K w_k \cdot \log \hat{\psi}_k}{\sum_{k=1}^K w_k} \right) \quad (3)$$

- ▶ Study index: $k = 1, \dots, K$
- ▶ Weights: $w_k = 1 / \widehat{\text{Var}}(\log \hat{\psi}_k)$ (\rightarrow fixed effect model)
- ▶ See formulae (1) and (2) for definition of $\hat{\psi}_k$ and $\widehat{\text{Var}}(\log \hat{\psi}_k)$
- ▶ Analogous for risk ratio as effect measure: $\log \hat{\phi}_k$
- ▶ For risk difference: $\hat{\eta}_k$ (without exp function in equation (3))

Meta-Analysis of CR – Inverse Variance Method

```
m <- metabin(crHDCT, nHDCT, crControl, nControl,
             data = cr, studlab = study,
             sm = "OR", method = "Inverse", comb.random = FALSE)
summary(m)

## Number of studies combined: k=14
##
##              OR              95%-CI          z p-value
## Fixed effect model 1.3228 [1.0999; 1.5909] 2.9713  0.003
##
## Quantifying heterogeneity:
## tau^2 = 0.0897; H = 1.3 [1; 1.78]; I^2 = 41% [0%; 68.6%]
##
## Test of heterogeneity:
##      Q d.f.  p-value
## 22.03  13  0.0549
##
## Details on meta-analytical method:
## - Inverse variance method
```

Meta-Analysis of CR – Inverse Variance Method

```
m4 <- rma(ai = crHDCt, n1i = nHDCt, ci = crControl, n2i = nControl,
  data = cr, measure = "OR", method = "FE")

m4

##
## Fixed-Effects Model (k = 14)
##
## Test for Heterogeneity:
## Q(df = 13) = 22.0277, p-val = 0.0549
##
## Model Results:
##
## estimate      se      zval      pval      ci.lb      ci.ub      **
## 0.2798      0.0942      2.9713      0.0030      0.0952      0.4643
##
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

Mantel-Haenszel Method – Odds ratio – Definition

Mantel and Haenszel (1959), JNCI:

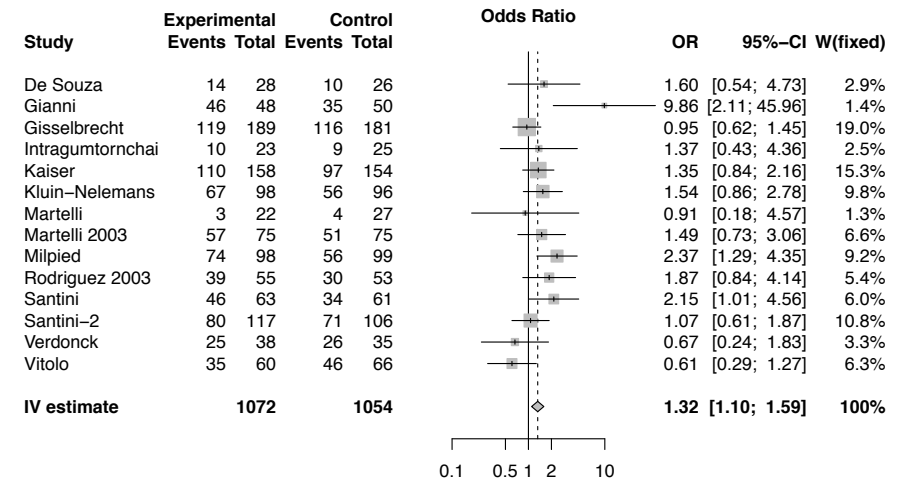
- ▶ Estimator for common odds ratio in stratified case-control study
- ▶ Can be used in meta-analysis of RCTs
- ▶ Fixed effect method

Mantel-Haenszel odds ratio $\hat{\psi}_{MH}$:

$$\hat{\psi}_{MH} = \frac{\sum_{k=1}^k w_k \cdot \hat{\psi}_k}{\sum_{k=1}^k w_k} \quad (4)$$

- ▶ Weights: $w_k = \frac{b_k c_k}{n_k}$

Forest Plot – CR – Inverse Variance Method



Meta-Analysis of CR – Mantel-Haenszel Method

```
m.mh <- update(m, method = "MH")
summary(m.mh)

## Number of studies combined: k=14
##
## OR      95%-CI      z      p-value
## Fixed effect model 1.3459 [1.1226; 1.6137] 3.2093 0.0013
##
## Quantifying heterogeneity:
## tau^2 = 0.0897; H = 1.3 [1; 1.78]; I^2 = 41% [0%; 68.6%]
##
## Test of heterogeneity:
## Q d.f. p-value
## 22.03 13 0.0549
##
## Details on meta-analytical method:
## - Mantel-Haenszel method
```

Meta-Analysis of CR – Mantel-Haenszel Method

```
rma.mh(ai = crHDCt, n1i = nHDCt, ci = crControl, n2i = nControl,
       data = cr, measure = "OR")

##
## Fixed-Effects Model (k = 14)
##
## Test for Heterogeneity:
## Q(df = 13) = 22.0615, p-val = 0.0544
##
## Model Results (log scale):
##
## estimate      se      zval      pval      ci.lb      ci.ub
## 0.2971 0.0926 3.2093 0.0013 0.1157 0.4785
##
## Model Results (OR scale):
##
## estimate      ci.lb      ci.ub
## 1.3459 1.1226 1.6137
##
## Cochran-Mantel-Haenszel Test: CMH = 10.0612, df = 1, p-val = 0.0015
```

Peto Odds Ratio (Yusuf et al., 1985)

Peto Odds Ratio ψ^* :

$$\hat{\psi}^* = \exp\left(\frac{a - E(a|\dots; \psi = 1)}{\text{Var}(a|\dots; \psi = 1)}\right) \quad (5)$$

with

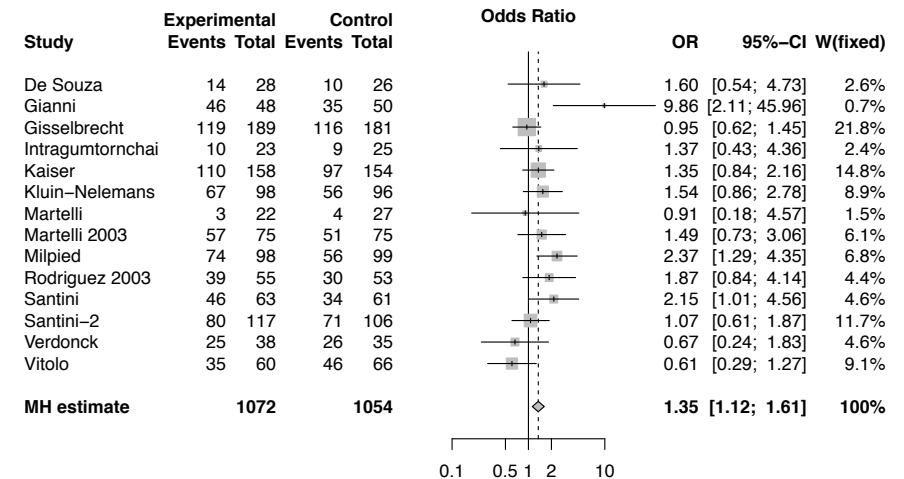
- ▶ Four fixed marginal totals: '...'
- ▶ Expected cell count:

$$E(a|\dots; \psi = 1) = \frac{(a+b)(a+c)}{n}$$

- ▶ Hypergeometric variance of cell count a :

$$\text{Var}(a|\dots; \psi = 1) = (a+b)(c+d)(a+c)(b+d)/(n^2(n-1)) \quad (6)$$

Forest Plot – CR – Mantel-Haenszel Method



Peto Method – Definition

Yusuf et al. (1985):

- ▶ Variant of the inverse variance method using Peto odds ratio and its variance
→ Dedicated method for odds ratio as summary measure
- ▶ Fixed effect method

Overall Peto odds ratio $\hat{\psi}_{Peto}$:

$$\hat{\psi}_{Peto} = \exp\left(\frac{\sum_{i=1}^k w_i^* \cdot \log \hat{\psi}_i^*}{\sum_{i=1}^k w_i^*}\right) \quad (7)$$

- ▶ Weights: $w_i^* = 1/\widehat{\text{Var}}(\log \hat{\psi}_i^*)$
- ▶ See formulae (5) and (6) for definition of $\hat{\psi}_i^*$ and $\widehat{\text{Var}}(\log \hat{\psi}_i^*) = 1/\text{Var}(a|\dots; \psi = 1)$

Effect measure	Estimate	95%-CI
Risk ratio $\hat{\psi}_{IV}$	1.1157	[1.0493; 1.1864]
Risk ratio $\hat{\psi}_{MH}$	1.1076	[1.0404; 1.1791]
Odds ratio $\hat{\phi}_{IV}$	1.3228	[1.0999; 1.5909]
Odds ratio $\hat{\phi}_{MH}$	1.3459	[1.1226; 1.6137]
Odds ratio $\hat{\phi}_{Peto}$	1.3462	[1.1233; 1.6134]
Risk difference $\hat{\eta}_{IV}$	0.0715	[0.0325; 0.1105]
Risk difference $\hat{\eta}_{MH}$	0.0656	[0.0261; 0.1051]

Peto method:

- ▶ R function **metabin**, argument `method = "Peto"`
- ▶ R function **rma.peto**

Random effects estimate $\hat{\psi}_{RE}$ (Fleiss, 1993):

$$\hat{\psi}_{RE} = \exp \left(\frac{\sum_{k=1}^K w_k^* \cdot \log \hat{\psi}_k}{\sum_{k=1}^K w_k^*} \right)$$

- ▶ Study index: $k = 1, \dots, K$
- ▶ Weights: $w_k^* = 1 / (\widehat{\text{Var}}(\log \hat{\psi}_k) + \hat{\tau}^2)$ (→ random effects model)
- ▶ See Session 1 for estimation of between-study variance τ^2
- ▶ Calculated in addition to fixed effect estimate by default in R function **metabin** (see arguments `comb.random` and `method.tau`)
- ▶ Default in R function **rma.uni** (see argument `method`)

Availability of methods:

Method	OR	RR	RD	other
Inverse Variance	×	×	×	×
Mantel-Haenszel	×	×	×	–
Peto	×	–	–	–

Properties for binary outcomes:

- ▶ Inverse variance method performs poor in meta-analyses with small studies
- ▶ Peto method performs poor in unbalanced designs and nearly balanced designs if odds ratio differs substantially from 1.00 (Greenland and Salvani, 1990)
- ▶ Peto method performs well in meta-analysis with very sparse data (Bradburn et al., 2007)
- ▶ MH approach recommended as method of choice (Emerson, 1994)

- ▶ Fixed effect model:
Inverse variance method inferior to Mantel-Haenszel and Peto method
- ▶ Fixed effect model often not reasonable
→ Random effects model (based on inverse variance method)
- ▶ Problems of inverse variance method (Stijnen et al., 2010):
 1. Variance estimate $\widehat{\text{Var}}(\log \hat{\psi}_k)$ assumed to be known (uncertainty not taken in account)
 2. Normal distribution assumption for $\log \hat{\psi}_k$ might not be justified
 3. (!) $\log \hat{\psi}_k$ and $\widehat{\text{Var}}(\log \hat{\psi}_k)$ are typically correlated (not taken into account)
 4. Additional difficulties in sparse binary data
- ▶ Stijnen et al. (2010): Use of generalised linear mixed models

Classic random effects model (Normal-Normal model):

$$\theta_k \sim N(\theta, \tau^2)$$

$$\hat{\theta}_k \sim N(\theta_k, \text{Var}(\hat{\theta}_k))$$

GLLM – Hypergeometric-Normal model:

- ▶ Model for odds ratio as effect measure
- ▶ Conditional on total number of events

$$\theta_k \sim N(\theta, \tau^2)$$

$$\hat{\theta}_k \sim \text{Non-central Hypergeometric (with argument } \theta_k)$$

GLLM – Hypergeometric-Normal model:

```
glmm1 <- rma.glmm(ai = crHDCt, n1i = nHDCt,
                 ci = crControl, n2i = nControl,
                 data = cr, measure = "OR",
                 model = "CM.EL")
```

`model = "CM.EL"`: conditional model with exact likelihood

GLLM – Binomial-Normal model:

```
glmm2 <- update(glmm1, model = "CM.AL")
```

`model = "CM.AL"`: conditional model with approximate likelihood

GLLM – Binomial-Normal model:

- ▶ Approximation to Hypergeometric-Normal model
- ▶ Applicable if total number of events is small relative to group sizes
- ▶ Number of events in experimental group a_{Tk} and control group c_{Tk} :

$$a_{Tk} \sim \text{Binomial}(a_{Tk} + c_{Tk}, p_k)$$

$$p_k = \frac{\exp(\log(n_{Tk}/n_{Ck}) + \theta_k)}{1 + \exp(\log(n_{Tk}/n_{Ck}) + \theta_k)}$$

with n_{Tk}, n_{Ck} number of patients in treatment groups

- ▶ Random intercept logistic regression model with offset $\log(n_{Tk}/n_{Ck})$

```
glmm1

##
## Random-Effects Model (k = 14; tau^2 estimator: ML)
## Model Type: Conditional Model with Exact Likelihood
##
## tau^2 (estimated amount of total heterogeneity): 0.0791 (SE = 0.0910)
## tau (square root of estimated tau^2 value):      0.2812
## I^2 (total heterogeneity / total variability):    37.99%
## H^2 (total variability / sampling variability):   1.61
##
## Tests for Heterogeneity:
## Wld(df = 13) = 21.8322, p-val = 0.0580
## LRT(df = 13) = 24.8475, p-val = 0.0242
##
## Model Results:
##
## estimate      se      zval      pval      ci.lb      ci.ub
## 0.3312      0.1274    2.5998    0.0093    0.0815    0.5810    **
##
```


GLMM - Results - Approximate Model

```
glmm2

##
## Random-Effects Model (k = 14; tau^2 estimator: ML)
## Model Type: Conditional Model with Approximate Likelihood
##
## tau^2 (estimated amount of total heterogeneity): 0
## tau (square root of estimated tau^2 value):      0
## I^2 (total heterogeneity / total variability):    0.00%
## H^2 (total variability / sampling variability):    1.00
##
## Tests for Heterogeneity:
## Wld(df = 13) = 6.4477, p-val = 0.9283
## LRT(df = 13) = 6.4827, p-val = 0.9268
##
## Model Results:
##
## estimate      se      zval      pval      ci.lb      ci.ub
## 0.1022      0.0542    1.8850    0.0594   -0.0041     0.2085
##
```

Summary

Meta-analysis with binary outcome

- ▶ Fixed effect model
 - ▶ Well established methods long available
- ▶ Random effects model:
 - ▶ Generalised linear mixed model preferable over inverse variance method
 - ▶ Exact method (Hypergeometric-Normal model) typically computational feasible in meta-analysis setting
 - ▶ Disadvantage of GLMMs: no forest plot

Comparison of results

```
# Classic random effects model (Normal-Normal model)
predict(update(m4, method = "ML"), transf = exp)

##      pred ci.lb ci.ub cr.lb cr.ub
## 1.3550 1.0788 1.7019 0.8337 2.2023

# GLMM - exact model (Hypergeometric-Normal model)
predict(glmm1, transf = exp)

##      pred ci.lb ci.ub cr.lb cr.ub
## 1.3927 1.0849 1.7877 0.7604 2.5507

# GLMM - approximate model (Binomial-Normal model)
predict(glmm2, transf = exp)

##      pred ci.lb ci.ub cr.lb cr.ub
## 1.1076 0.9959 1.2319 0.9959 1.2319
```

References

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Mantel-Haenszel Method – Risk ratio – Definition

Mantel-Haenszel risk ratio $\hat{\phi}_{MH}$:

$$\hat{\phi}_{MH} = \frac{\sum_{k=1}^K w_k \cdot \hat{\phi}_k}{\sum_{k=1}^K w_k}$$

- Weights: $w_k = \frac{(a_k + b_k)c_k}{n_k}$

Mantel-Haenszel Method – Odds ratio – Confidence int.

Robins et al. (1986a,b):

$$\widehat{\text{Var}}(\log \hat{\psi}_{MH}) = \frac{\sum_{k=1}^K P_k R_k}{2 \left(\sum_{k=1}^K R_k \right)^2} + \frac{\sum_{k=1}^K (P_k S_k + Q_k R_k)}{2 \sum_{k=1}^K R_k \sum_{k=1}^K S_k} + \frac{\sum_{k=1}^K Q_k S_k}{2 \left(\sum_{k=1}^K S_k \right)^2}$$

$$\text{with } P_k = \frac{a_k + d_k}{n_k}, Q_k = \frac{b_k + c_k}{n_k}, R_k = \frac{a_k d_k}{n_k}, \text{ and } S_k = \frac{b_k c_k}{n_k}$$

- Variance estimator robust both in sparse data and large strata models

- $(1 - \alpha)$ -confidence interval:

$$\exp(\log \hat{\psi}_{MH} \pm z_{1-\frac{\alpha}{2}} \text{S.E.}(\log \hat{\psi}_{MH}))$$

- Standard error $\text{S.E.}(\log \hat{\psi}_{MH}) = \sqrt{\widehat{\text{Var}}(\log \hat{\psi}_{MH})}$

Mantel-Haenszel Method – Risk ratio – Conf. int.

Greenland and Robins (1985):

$$\widehat{\text{Var}}(\log \hat{\phi}_{MH}) = \frac{\sum_{k=1}^K \frac{(a_k + b_k)(c_k + d_k)(a_k + c_k) - a_k c_k n_k}{n_k^2}}{\sum_{k=1}^K \frac{a_k(c_k + d_k)}{n_k} \sum_{k=1}^K \frac{c_k(a_k + b_k)}{n_k}}$$

- Robust variance estimator

- $(1 - \alpha)$ -confidence interval:

$$\exp(\log \hat{\phi}_{MH} \pm z_{1-\frac{\alpha}{2}} \text{S.E.}(\log \hat{\phi}_{MH}))$$

- Standard error $\text{S.E.}(\log \hat{\phi}_{MH}) = \sqrt{\widehat{\text{Var}}(\log \hat{\phi}_{MH})}$

Mantel-Haenszel Method – Risk difference – Definition

Mantel-Haenszel risk difference $\hat{\eta}_{MH}$:

$$\hat{\eta}_{MH} = \frac{\sum_{k=1}^K w_k \cdot \hat{\eta}_k}{\sum_{k=1}^K w_k}$$

► Weights: $w_k = \frac{(a_k + b_k)(c_k + d_k)}{n_k}$

Peto Method – Confidence interval

► Large sample variance estimate for logarithm of $\hat{\psi}_{Peto}$:

$$\widehat{\text{Var}}(\log \hat{\psi}_{Peto}) = \frac{1}{1 / \sum_{k=1}^K \widehat{\text{Var}}(\log \hat{\psi}_k^*)}$$

► $(1 - \alpha)$ -confidence interval:

$$\exp\left(\log \hat{\psi}_{Peto} \pm z_{1-\frac{\alpha}{2}} \text{S.E.}(\log \hat{\psi}_{Peto})\right)$$

► Standard error $\text{S.E.}(\log \hat{\psi}_{Peto}) = \sqrt{\widehat{\text{Var}}(\log \hat{\psi}_{Peto})}$

Mantel-Haenszel Method – Risk difference – Conf. int.

Greenland and Robins (1985):

$$\widehat{\text{Var}}(\hat{\eta}_{MH}) = \frac{\sum_{k=1}^K \frac{(a_k b_k n_{Ck})^3 + (c_k d_k n_{Tk})^3}{(n_{Tk} n_{Ck} (n_{Tk} + n_{Ck}))^2}}{\left(\sum_{k=1}^K \frac{(a_k + b_k)(c_k + d_k)}{n_k} \right)^2}.$$

► Robust variance estimator

► $(1 - \alpha)$ -confidence interval:

$$\hat{\eta}_{MH} \pm z_{1-\frac{\alpha}{2}} \text{S.E.}(\hat{\eta}_{MH})$$

► Standard error $\text{S.E.}(\hat{\eta}_{MH}) = \sqrt{\widehat{\text{Var}}(\hat{\eta}_{MH})}$